Learning Bayesian networks given a data set consisting of samples that are not independent and identically distributed ABNMS 2011

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- Consider a random vector $\boldsymbol{X} = (X_1, \dots, X_p)^T \sim N_p(\boldsymbol{0}, \Sigma)$
- Learning the structure of the Bayesian network of X usually requires n iid samples
- What if we have a more complex data set?
 - non-independent samples;
 - additional components of variance;
 - data on exogenous variables thought to affect X.

Learning Graphical Structure

- A Bayesian network B = (G, Θ) for a random vector X consists of two components:
 - a directed acyclic graph $G = (V, E), V = \{1, 2, \dots, p\},\$
 - conditional densities for each random variable, f(x_i|**x**_{Pi}, θ_i), where P_i is the set of parents of *i* in G, θ_i the parameters.

The graph and conditional densities specify a joint density function for *X*:

$$f(\mathbf{x}|\Theta) = \prod_{i=1}^{p} f(x_i | \mathbf{x}_{P_i}, \theta_i).$$

- Want to learn *G* given a data set $d = \{x_1, x_2, \dots, x_p\}, x_i = (x_{i1}, \dots, x_{in})$
 - A popular approach for learning about genetic regulatory networks.

- How well G describes the data is quantified by a score metric, S(G|d).
 - Score we consider is:

$$S(G|d) = p(G)p(d|G) = p(G) \int p(d|G,\Theta)p(\Theta|G)d\Theta.$$

The BGe score of Geiger and Heckerman (1994).

- Need to specify:
 - *p*(*G*): prior on space of DAGs;
 - $p(d|G,\Theta) = \prod_{i=1}^{p} f(\mathbf{x}_i | \mathbf{x}_{P_i}, \theta_i);$
 - $p(\Theta|G)$: prior for the parameters.

IID samples

When *d* consists of iid samples:

$$\boldsymbol{x}_i | \boldsymbol{x}_{P_i}, \gamma_i, \psi_i \sim N_n(\boldsymbol{x}_{P_i}\gamma_i, \psi_i I_n).$$

Parameter priors:

$$\begin{array}{ll} \gamma_i | \psi_i & \sim \textit{N}_{|\textit{P}_i|} \left(0, \frac{\psi_i}{\tau} \textit{I} \right), \\ \psi_i^{-1} & \sim \textit{Ga} \left(\frac{\delta + |\textit{P}_i|}{2}, \frac{\tau}{2} \right). \end{array}$$

These priors chosen to give an **equivalent** score, Geiger and Heckerman (2002).

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$$S(G|d) = \rho(G) \prod_{i=1}^{p} f(\boldsymbol{x}_{i}|\boldsymbol{x}_{P_{i}}),$$

$$f(\boldsymbol{x}_{i}|\boldsymbol{x}_{P_{i}}) = \int_{\mathbb{R}^{|P_{i}|} \times (0,\infty)} f(\boldsymbol{x}_{i}|\boldsymbol{x}_{P_{i}},\gamma_{i},\psi_{i}) f(\gamma_{i},\psi_{i}) d\gamma_{i} d\psi_{i}.$$

S(G|d) used in conjunction with algorithms for exploring the DAG space.

- Greedy hill climbing,
- High-dimensional Bayesian covariance selection, Dobra *et al.* (2004)

Grape gene example:

- Learn about the relationships of grape heat shock genes
- Grapes sampled from 3 geographically distinct vineyards
- Temperatures at times leading up to picking of grapes available.

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Must account for effects of exogenous variables!

Now have

$$\boldsymbol{x}_i | \boldsymbol{x}_{P_i}, \gamma_i, \boldsymbol{b}_i, \psi_i \sim N_n(\boldsymbol{x}_{P_i}\gamma_i + Q\boldsymbol{b}_i, \psi_i \boldsymbol{I}_n)$$

where

$$Q = (\boldsymbol{q}_1 | \cdots | \boldsymbol{q}_m).$$

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$$egin{aligned} b_i | \psi_i &\sim \mathcal{N}_m(0, v^{-1}\psi_i I) \ S_\mathcal{B}(G|d) &= p(G) \prod_{i=1}^p f_v(oldsymbol{x}_i | oldsymbol{x}_{\mathcal{P}_i}). \end{aligned}$$

 Residual approach: remove random effects by analysing residuals: n × (n − m) matrix P:

$$P^T Q = 0$$
, $P^T P = I_{n-m}$, $PP^T = I_n - Q(Q^T Q)^{-1}Q^T$.

Then

$$P^{T}\boldsymbol{x}_{i}|P^{T}\boldsymbol{x}_{P_{i}},\gamma_{i},\psi_{i}\sim N_{n-m}\left(P^{T}\boldsymbol{x}_{P_{i}}\gamma_{i},\psi_{i}I\right).$$

Using previous priors for γ_i and ψ_i , get score metric

$$S_R(G|d) = p(G) \prod_{i=1}^p f_R(P^T \boldsymbol{x}_i | P^T \boldsymbol{x}_{P_i}).$$

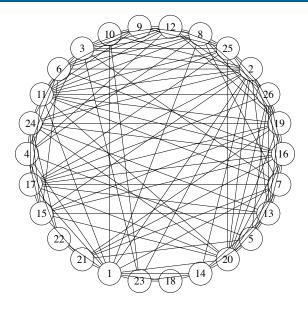
Grape Gene Example

- *n* = 50 samples of *p* = 26 grape berry genes;
- Grape berries sampled from 3 vineyards
- Genes are heat shock genes and we have temperature measurements.
- Assume the following model for each gene:

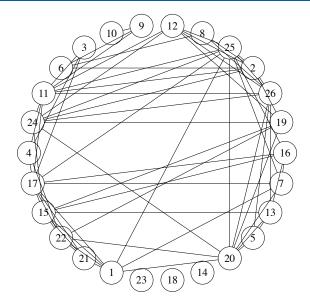
$$\begin{array}{lll} X_{ij} & = & \displaystyle\sum_{l \in \mathcal{P}_i} \gamma_{il} X_{lj} + \displaystyle\sum_{r=1}^m q_{rj} b_{ir} + \epsilon_{ij}, \ \epsilon_{ij} \sim \mathcal{N}(0,\psi_i), \\ \gamma_{il} & \sim & \mathcal{N}(0,\tau^{-1}\psi_i), \\ \psi_i & \sim & \operatorname{Inverse \ Gamma}\left(\frac{\delta + |\mathcal{P}_i|}{2}, \frac{\tau}{2}\right). \end{array}$$

Use the residual approach to account for b_i . (Is this a good idea?)

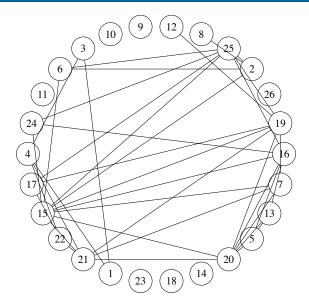
Ignoring effects of vineyard and temperature



Residual approach, vineyard effects



Residual approach, vineyard and temperature effects



- As more variation due to exogenous sources is accounted for, graphs become sparser
- Genes 14, 18, 23: disconnected from rest of graph in last two graphs
 - Expressions of these genes have very low ses no variation to be explained by relationships with other genes!
- Genes 9, 10, 11: correspond to HSP 81, early response to dehydration
 - Role not very well understood, our analysis indicates they are not implicated in heat shock gene network.

Comparing Bayesian and residual approaches

- Should we have used the Bayesian approach in the grape gene example?
 - Residual approach is easier to use;
 - May obtain less information about γ_i, ψ_i :
 - May be important for posterior estimation of parameters.

Comparing Bayesian and residual approaches

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 - Residual approach is easier to use;
 - May obtain less information about γ_i, ψ_i :
 - May be important for posterior estimation of parameters.
- Full Bayesian approach posterior: *f_B*(*γ_i*, *ψ_i*|*x_i*, *x_{P_i}). Residual approach posterior: <i>f_R*(*γ_i*, *ψ_i*|*x_i*, *x_{P_i}).*
- We consider the Kullback Leibler divergence:

$$D\{f_B, f_R\} = \int_{\mathbb{R}^{|P_i|}} \int_0^\infty f_B \log\left(\frac{f_B}{f_R}\right) d\psi_i d\gamma_i.$$

Divergence for marginal covariance matrix Σ

$$var(\boldsymbol{X}|\{\boldsymbol{\gamma}_i,\psi_i\}_{i=1,\dots,p})=\boldsymbol{\Sigma}$$

• If the true graphical structure of **X** is known:

$$D_{\Sigma} \{ f_{B}(\Sigma | \boldsymbol{X}), f_{R}(\Sigma | \boldsymbol{X}) \}$$

= $\sum_{i=1}^{p} D \{ f_{B}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i}, \boldsymbol{x}_{P_{i}}), f_{R}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i}, \boldsymbol{x}_{P_{i}}) \}.$

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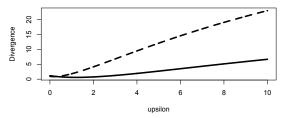
• Divergence for the empty graph:

$$D_{\Sigma}^{e} = \sum_{i=1}^{p} D\left\{f_{B}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i}), f_{R}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i})\right\},\$$

• Divergence for an arbitrary full graph:

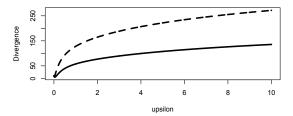
$$D_{\Sigma}^{f} = \sum_{i=1}^{p} D\left\{f_{B}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i}, \boldsymbol{x}_{1}, \dots, \boldsymbol{x}_{i-1}), f_{R}(\boldsymbol{\gamma}_{i}, \psi_{i} | \boldsymbol{x}_{i}, \boldsymbol{x}_{1}, \dots, \boldsymbol{x}_{i-1})\right\}.$$

Divergence in Grape Gene Example



Vineyards





- The Bayesian and residual score metrics extend the utility of score-based methods for learning networks to situations where the data does not consist of iid samples.
- Provided sample size is not too small, residual approach is a useful alternative to Bayesian approach
 - Even when the assumptions of the Bayesian approach are valid.

Some questions:

- What happens when the chosen prior distribution of the effects of exogenous variables is not suitable?
- Are there situations where the residual approach performs better than the Bayesian approach?